# Abstract

**Aims:**  To develop and externally validate a parsimonious statistical prediction model of 90-day mortality after elective Total Hip Arthroplasty (THA), and to provide a web-calculator for clinical usage.

**Patients and methods:**  We included 53,099 patients with cemented THA due to osteoarthritis from the Swedish Hip Arthroplasty Register for model derivation and internal validation, as well as 125,428 patients from the National Joint Register from England and Wales for external model validation. A model was developed using a bootstrap ranking procedure with a least absolute shrinkage and selection operator (LASSO) logistic regression model combined with piecewise linear regression. Discriminative ability was evaluated by the area under the receiver operating characteristics curve (AUC). Calibration belt plots were used to assess model calibration.

**Results:**  A main effects model combining age, sex, American Society for Anaesthesiologists (ASA) class, the presence of cancer, diseases of the central nervous system, kidney disease, and diagnosed obesity had good discrimination, both internally (AUC = 0.78, 95 % CI 0.75 - 0.81) and externally (AUC = 0.75, 95 % CI 0.73 - 0.76). This model was superior to traditional models based on the Charlson (AUC = 0.66, 95 % CI 0.62 - 0.70) and Elixhauser (AUC = 0.64, 95 % CI 0.59 - 0.68) comorbidity indices. The model was well calibrated for predicted probabilities up to 5%.

**Conclusion:**  We developed a parsimonious model which may facilitate individualized risk assessment prior to one of the most common surgical interventions. We published a web-calculator to aid clinical decision-making (<https://shpr.registercentrum.se/om-registret-1/forskning/prediktion-av-90-dagarsmortalitet/p/SkyeTsTFB>).

**Clinical relevance:**

* We developed a parsimonious model for risk assessment of 90-day mortality after elective THA.
* The model considers age, sex, ASA class, the presence of cancer, diseases of the central nervous system, kidney disease, and diagnosed obesity.
* We published an on-line web-calculator (<https://shpr.registercentrum.se/om-registret-1/forskning/prediktion-av-90-dagarsmortalitet/p/SkyeTsTFB>).

##### PAGE BREAK

# Introduction

Shared decision making has evolved into an integral part of patient-physician interactions prior to surgical interventions, and the weighing of risks against benefits is central to this process. 90-day mortality after THA performed due to osteoarthritis is low, ranging from 0.2 % to 0.6 %.1 The risk of short-term mortality is nevertheless an essential part of preoperative discussions between patients and surgeons.

Comorbidity is associated with a shorter remaining life span, but the Charlson and Elixhauser comorbidity indices poorly predict mortality after THA.1,2,2–5 Additionally, these complex comorbidity instruments are based on information from extensive in- or out-patient databases, each measure being defined by 1,178 and 1,516 individual International Classification of Diseases (ICD)-10 codes, respectively. The use of these indices is therefore limited to research settings.

Pre-operative comorbidity data have been used for prediction models of early mortality and adverse events such as surgical site infection, in the context of THA surgery.6–12 The first prediction model based on a European cohort was developed in the UK (<https://jointcalc.shef.ac.uk/>). However, it remains unclear to what level the depth of comorbidity characterization influences the accuracy of mortality prediction models.

We aimed to develop a parsimonious prediction model of 90-day mortality after THA with internal and external validation of discrimination and calibration, and to compare this to the accuracy of prediction models based on existing comorbidity measures.13

# Participants and Methods

We used data from the Swedish Hip Arthroplasty Register (SHAR) for model derivation and internal validation. The model was then validated externally on patients from England and Wales recorded in the National Joint Registry for England, Wales, Northern Ireland, the Isle of Man and the States of Guernsey (NJR). We focused on patients with cemented THA due to osteoarthritis since cementation is the most common fixation technique used in Sweden, and we thus excluded patients where either the femoral stem and/or the acetabular cup were cementless (cementless, hybrids, and reverse hybrids). Only the last operated hip was accounted for in patients with bilateral THA (Figure 1).14

## Derivation cohort (Sweden)

All public and private Swedish hospitals report their THAs to SHAR, yielding a coverage of 100 %, with an individual patient completeness of 98 %.15 Our inclusion period started 2008 since the American Society for Anaesthesiologists (ASA) class and Body Mass Index (BMI) were systematically recorded in SHAR from then on. The observation period ended 2015, since we only had access to information on comorbidity from the national patient register up to this point. Deterministic data linkage was performed by 10-digit identity numbers, assigned to all Swedish residents at birth or immigration, which has been described previously.16 Age, sex, BMI, ASA class, type of hospital and year of surgery were collected from the SHAR. Data on educational level were recorded for more than 98 % of the population with 85 % accuracy in the longitudinal integration database for health insurance and labour market studies (LISA) from Statistics Sweden.17 Civil status, were also collected from LISA and comorbidity during the year preceding index surgery from the National Patient Register (NPR). NPR contains diagnoses coded by ICD-10 from in- and outpatient episodes in all private and public hospitals. Completeness for NPR is above 99 % and 85-95 % of all diagnostic codes are valid.18 Death dates were linked from the national population register.

Patients with missing information on BMI or a measurement above 50 were excluded, as were patients with missing information on ASA class or class IV and above, as well as patients with unknown educational level or type of hospital. 53,099 patients (mean age 73 years, range 35 - 99 and 61 % females) were included (Figure 1).

## External validation cohort (England and Wales)

Similar inclusion criteria were applied to the validation cohort. We excluded no patients based on missing information on BMI, educational level, or civil status however, since those variables were not used in the final model. 125,428 patients were included (Figure 2).

Reporting to NJR is mandatory for all hospitals whether run by the National Health Service (NHS) or privately in England and Wales. Patient participation is based on informed consent. Consent rates vary slightly between years and regions but it was 92.3 % in both England and Wales in 2018.19,20 NJR data was linked to the secondary uses service database from the National Health Service (NHS) with comorbidity from the hospital episode statistics (HES) database in England, and the patient episode database for Wales.21 Out-patient deaths were linked from the office of national statistics based on NHS-numbers, as well as patient name, age, sex and address.22

## Defining co-morbidity

Comorbidity was defined by individual ICD-10 codes grouped into 17 categories by Charlson and 31 by Elixhauser.23 Some conditions were identified by both indices, and some were closely related and therefore merged (such as hypertension with and without complications, or abuse of either drugs or alcohol). We combined individual diagnostic groups to establish 21 broader categories of comorbidity (Table 1).

## Statistical analysis for model development

We used a statistical procedure with bootstrap ranking and a logistic least absolute shrinkage and selection operator (LASSO; figure 3 and 4).24 Age and BMI were normalized to have mean = 0 and standard deviation = 1. Comorbidities recorded for at least one patient who died within 90 days, and one who did not, were included in the modelling process. One hundred bootstrap samples were drawn from the observed data set. We used ten-fold cross validation for every bootstrap sample with a range of potential penalty values (:s) in a logistic LASSO model. We kept :s minimizing the mean cross-validated deviances and those were used to estimate model coefficients for each potential predictor. The means of the magnitude (absolute values) of those estimates were used as measures of variable importance. Piece-wise linear regression was used to detect a breaking point where a significant drop in variable importance was observed. Potential predictors with variable importance above this breaking point were kept as model candidates. The whole process was repeated 100 times. Covariates that were selected at least once were used in a main effects model of multivariable logistic regression without penalty, and without pre-normalization of numeric variables (main model). A reduced model with variables chosen at least 33 out of the 100 times were used as a simpler alternative. This model was also evaluated without cancer as a predictor, considering that medical indications for THA surgery may be different for patients with cancer or without cancer. Univariable models with the ASA class, as well as the Charlson or Elixhauser comorbidity indices were used for benchmarking, as well as a model including only age and sex. Each model including age was fitted three times, once with age as a main effect and twice with restricted cubic splines, either by two or three knots. Odds ratios for the final model were estimated with 95 % confidence intervals and -values below 0.05 were considered statistically significant.

## Statistical analysis for model validation

Each model was used to predict the probability of death within 90 days for patients from the SHAR (internal validation). Receiver operating characteristic (ROC) curves were drawn and the area under those curves (AUC) were used as a measure of discriminative ability. Models with a lower bootstrap-based 95 % confidence limit above 0.7, were considered good. We used the bias-corrected Somers’ rank correlation based on 100 resamples to adjust for optimism.25 Calibration bands were made to graphically assess model calibration, comparing predicted probabilities and observed proportions.26 The reduced model was then evaluated externally. An AUC with 95 % CI was calculated for the model as-is. Re-calibration of the model intercept was then performed to account for different mortality rates in Sweden compared to England and Wales. An updated over-all slope was also calculated to account for country-specific treatment differences.27

## Statistical tools

We built an online web calculator (<https://shpr.registercentrum.se/om-registret-1/forskning/prediktion-av-90-dagarsmortalitet/p/SkyeTsTFB>) to be used in clinical practice.

We used R version 3.6.1 (R Foundation for Statistical Computing, Vienna, Austria) with significant packages tidyverse, tidymodels, furrr, pROC and shiny. All R-scripts and necessary configurations (but no personal data) are available at <http://doi.org/10.5281/zenodo.3732852>. A linked Binder environment is also available for interactive usage.

## Ethical approval

Ethical approval for this study was obtained from the Regional Ethical Review Board in Gothenburg (360-13) and by the NJR in the UK (RSC2017/21). Informed consent was not mandatory according to the Swedish patient data law from 2009, and the UK law for pseudonymised data.

# Results

## Study participants

175 [0.33 % (95 % CI 0.28 - 0.38)] of the patients from SHAR died within 90 days and no one was censored before that. The median age of patients who died was 79 years (range 57-95; Table 2). Patients from the NJR had an unadjusted risk of 90-day mortality of 0.52 % (95 % CI 0.49 - 0.56). Further cohort characteristics are presented in Table 3.

## Model development and internal validation

There were five comorbidities not recorded for any patient who died in the Swedish derivation cohort and therefore not considered as potential predictors: acquired immunodeficiency syndrome by the human immunodeficiency virus (AIDS/HIV), coagulopathy, fluid electrolyte disorders, liver disease, and weight loss. The derived main model included age, sex, ASA class, the presence of cancer, diseases of the central nervous system, kidney disease, diagnosed obesity, heart condition, anaemia, and myocardial infarction. The reduced model was restricted to age, sex, ASA class, the presence of cancer, diseases of the central nervous system, kidney disease, and diagnosed obesity (Table 4).

There were no differences between models including age as a main effect or as restricted cubic splines. We therefore focused on the more parsimonious models with age as a main effect. Similarly, the correction for optimism only affected the third decimals of the AUC confidence intervals and was therefore omitted.

The main and reduced models were no different regarding discriminative ability (AUC = 0.79, 95 % CI 0.75 - 0.82 versus AUC = 0.78, 95 % CI 0.75 - 0.81). We therefore considered the reduced model as superior due to its simplicity. Traditional models performed poorly with 95 % confidence intervals not above 0.7: The Charlson comorbidity index had an AUC of (AUC = 0.66, 95 % CI 0.62 - 0.70) and the Elixhauser comorbidity an AUC of (AUC = 0.64, 95 % CI 0.59 - 0.68; Figure 5 and Figure 7).

The ability of the reduced model to estimate probabilities of death within 90 days is further illustrated in Figure 8 and 9. Model calibration was good for estimated probabilities up to 3 % and acceptable up to 5 % (Figure 10). Estimated model coefficients and corresponding odds ratios for the reduced model are presented in Table 5. Omitting cancer from the reduced model did not affect the AUC or calibration for estimated probabilities below 3 %, but calibration outside this range deteriorated, and we thus retained cancer as an important predictor.

## External validation

The discriminative ability for the reduced model was not statistically significantly different when applied to the external validation cohort (AUC = 0.75, 95 % CI 0.73 - 0.76) compared to the internal validation (AUC = 0.78, 95 % CI 0.75 - 0.81) (Figure 6 and Figure 7). Calibration of the re-calibrated model was slightly inferior compared to the internal calibration. Predicted probabilities between 0.5 % and 1.5 % were lower than observed proportions and the estimated 95 % confidence bands were wider. Over-all, calibration was good for predicted probabilities below 3 % and acceptable below 5 % (Figure 10).

# Discussion

We found that a multivariable main effects logistic regression model with age, sex, ASA class, the presence of cancer, diseases of the central nervous system, kidney disease, and diagnosed obesity was better at predicting death within 90 days after THA than traditional models based on the Charlson or Elixhauser comorbidity indices. The resulting model predicted the probability of death within 90 days as: where is the vector of estimated coefficients (including intercept) presented in table 5. This formula is considered valid for patients aged 57-95 years and for predicted probabilities up to 5 %.

## Clinical usage

Our model could be used in clinical practice, either by the formula, or by the web calculator (<https://shpr.registercentrum.se/om-registret-1/forskning/prediktion-av-90-dagarsmortalitet/p/SkyeTsTFB>). A 57-year-old woman with ASA class I and none of the important comorbidities would have a 0.17 ‰ risk to die within 90 days of surgery. A woman, 67 years old, would have an elevated risk of 0.37 ‰. A 95-year-old man with ASA class III and cancer, would have a risk of 6.5 %. Note, however, that covariate patterns with observed probabilities above 5 % were rare (0.13 %, n = 70). Risks above 5 % are therefore subject to extrapolation. The true proportions are likely lower (as indicated by Figure 10). Some risk calculators ignore this problem,6,12 but we think this should be acknowledged.

Variables in our model were chosen based on statistical properties and should therefore not be assigned any exact epidemiological and/or causal meaning per se. Nevertheless, age and sex are well-known predictors of remaining life span.28 ASA class III has the largest estimated coefficient among all predictors, indicating large relative importance. This is clinically reasonable since a label of “severe systemic disease” should be based on a relevant patient assessment prior to surgery.

Diagnosed obesity was not statistically significant ( 0.05) but was still relevant as a predictor due to unobserved heterogeneity. We noted that the proportion of patients in the SHAR with BMI above 30, the WHO definition of obesity, was much higher (23 %, n = 12,336) than the proportion of patients with a diagnosis code for obesity (ICD-10 = E66) recorded in the NPR (1.9 %, n = 1,000). BMI is systematically recorded in the SHAR with only around 5 % missing data, but the diagnosis of obesity in patients registered in the NJR might not have been made during previous contacts with the healthcare system.

It should be noted that short-term mortality after elective THA is rare and might therefore not be the most important issue for patients making an informed decision on whether to consent to THA surgery. Other important outcomes such as adverse events after surgery, including prosthetic joint infections (PJI) or dislocation, are much more common and have a considerable impact on quality of life and the necessity of subsequent surgeries. Our next step is therefore to extend the web calculator to adverse events such as PJI and dislocation.

## Strengths and limitations

The nationwide design with data from both Sweden, England and Wales is a strength. The Swedish registers are valid with low proportions of missing data.15,18 Some concerns have been raised regarding the HES database from UK, but a systematic review found that the overall median diagnostic accuracy (comparing ICD-codes to individual case notes) was 80 %.29 Both in- and outpatient comorbidity data were available for Sweden, but only in-patient data for England and Wales. The linkage between NJR and HES has been previously described,30 where it was noted that patients operated at privately funded units were not included in HES. 17 % of the patients had missing personal data or did not allow linkage, and 6 % were not found in HES although their data were available from the NJR.

A weakness of our study is that it is restricted to patients operated with cemented THA which impacts generalizability of our findings to countries with a much higher sage of cementless or hybrid techniques. Cementation is the most commonly used mode of THA fixation in Sweden, but not so in England and Wales where especially younger patients are more likely to be operated with cementless implants. To only include cemented THA therefore implies an over-representation of older and frailer patients from England and Wales and such patients are more likely to die within 90 days. This interpretation agrees with the calibration plot where estimated probabilities below 1.5 % underestimated the observed proportion of deaths in the external validation cohort.

We attempted at developing a similar model on a Swedish cohort of patients with uncemented, hybrid or reverse hybrid THAs. Unfortunately, it was impossible to analyze each mode of fixation separately since too few deaths occurred in each subgroup. It was possible to develop a model for all combined fixation modes other than totally cemented, but the 95 % CI obtained for the estimated AUC was very wide (0.54–0.76) and therefore of little relevance. Possible adaptations of our model to other settings or countries should therefore always be preceded by additional external validation with data from that target population.

We investigated whether the effect of age on mortality was non-linear using restricted cubic splines and a similar approach might be relevant for BMI. A U-shaped association between mortality and BMI categorized as underweight (below 18.5 kg/m2), normal (18.5 to 25 kg/m2), or overweight/obese (above 25 kg/m2) has been reported.31 Our recordings of BMI might be imprecise however. Some hospitals report actual measurements but it is unknown to what extent those data are self-estimated by the patients or by health care personnel. Therefore, we aimed to avoid overfitting BMI based on too elaborate transformations. To not include BMI treated as a continuous variable during external validation was, however, beneficial, since the completeness of this variable has traditionally been low in the NJR; 60 % in 2010 but since then increasing up to 80 % in 2018. Instead, the related ICD-10 code was used to capture diagnosed obesity from the Swedish NPR and the UK HES, respectively.

It should be noted that the risk model does not study THA as an observed intervention. We merely followed the cohort who already had THA. Deaths within 90 days might occur for those patients regardless if THA is inserted or not. The proximity in time however, the maximum of 90 days from THA to death, is an indication that the operation might be the main cause of death for the majority of deceased patients. This notion is supported by a study performed by Cnudde et al. who showed that patients selected for elective THA have a reduced one-year mortality risk when compared to the general population of the same age, sex, and year of birth. However, within the first 90 days after cemented THA surgery, the risk of death is increased when compared to the reference population, indicating that the surgery itself is associated with increased short-term mortality.1,32

## Conclusion

Our results indicate that the risk of early postoperative mortality after cemented THA can be pre-operatively assessed by a parsimonious prediction model. We hope that this model, with its accompanying web calculator, will facilitate shared decision-making.

# 

# Tables

Table 1: Categorization of comorbidities based on diagnostic groups according to the Charlson and Elixhauser comorbidity indices. Each main comorbidity category was first defined based on comorbidity data from the Swedish National patient register and from the Hospital Episodes statistics registry in England and Wales, coded along the International Classification of Diseases version 10 (ICD-10) and categorized according to Quan et al.23 (CNS = central nervous system.)

|  |  |  |
| --- | --- | --- |
| Comorbidities by groups | Charlson | Elixhauser |
| AIDS/HIV | AIDS/HIV | AIDS/HIV |
| Anaemia |  | Blood loss anaemia, Deficiency anaemia |
| Arrhythmia |  | Cardiac arrhythmias |
| Arterial hypertension |  | Hypertension uncomplicated, Hypertension complicated |
| Cancer | Malignancy, Metastatic solid tumour | Lymphoma, Metastatic cancer, Solid tumour |
| CNS disease | Dementia, Hemiplegia or paraplegia | Depression, Paralysis, Other neurological disorders, Psychoses |
| Coagulopathy |  | Coagulopathy |
| Diabetes | Diabetes without complication, Diabetes complication | Diabetes uncomplicated, Diabetes complicated |
| Drug alcohol abuse |  | Alcohol abuse, Drug abuse |
| Fluid electrolyte disorders |  | Fluid electrolyte disorders |
| Heart condition | Congestive heart failure | Congestive heart failure, Valvular disease |
| Myocardial infarction | Myocardial infarction |  |
| Hypothyroidism |  | Hypothyroidism |
| Kidney disease | Renal disease | Renal failure |
| Liver disease | Mild liver disease, Moderate or severe liver disease | Liver disease |
| Lung airways disease | Chronic pulmonary disease | Chronic pulmonary disease, Pulmonary circulation disorder |
| Diagnosed obesity |  | Obesity |
| Peptic ulcer | Peptic ulcer disease | Peptic ulcer disease |
| Rheumatic disease | Rheumatic disease | Rheumatoid arthritis |
| Vascular disease | Peripheral vascular disease, Cerebrovascular disease | Peripheral vascular disorder |
| Weight loss |  | Weight loss |

##### PAGE BREAK

Table 2: Age distribution among patients who died in the Swedish derivation cohort (range 57-95years).

|  |  |
| --- | --- |
| Age | prop |
| 35 - 65 | 0.10 % (10 of 9,633) |
| 66 - 70 | 0.22 % (25 of 11,271) |
| 71 - 75 | 0.22 % (28 of 12,530) |
| 76 - 80 | 0.37 % (40 of 10,788) |
| 81 - 85 | 0.56 % (36 of 6,486) |
| 86 - 90 | 1.45 % (31 of 2,139) |
| 91 - 99 | 1.98 % (5 of 252) |

##### PAGE BREAK

Table 3: Characteristics of the study population in the model derivation cohort (SHAR) and external validation cohort (NJR). Educational levels were classified as low (up to 9 years), middle (10-12 years) and high (at least 12 years). BMI = Body mass index. ASA = American Society for Anaesthesiologists classification. CNS = central nervous system.

|  |  |  |  |
| --- | --- | --- | --- |
| what | level | SHAR | NJR |
| n |  | 53,099 | 125,428 |
| Age (mean (SD)) |  | 72.68 (7.76) | 73.51 (8.67) |
| Sex = Female (%) |  | 32,440 (61.1) | 82,247 (65.6) |
| BMI (mean (SD)) |  | 27.19 (4.39) |  |
| ASA class (%) |  |  |  |
|  | I | 9,589 (18.1) | 10,890 (8.7) |
|  | II | 33,881 (63.8) | 89,388 (71.3) |
|  | III | 9,629 (18.1) | 25,150 (20.05) |
| Hospital (%) |  |  |  |
|  | University | 24,460 (46.1) |  |
|  | County | 16,507 (31.1) |  |
|  | Rural | 9,940 (18.7) |  |
|  | Private | 2,192 ( 4.1) |  |
| Education (%) |  |  |  |
|  | Low | 11,645 (21.9) |  |
|  | Middle | 20,604 (38.8) |  |
|  | High | 20,850 (39.3) |  |
| Civil status (%) |  |  |  |
|  | Married | 29,436 (55.4) |  |
|  | Single | 12,888 (24.3) |  |
|  | Widow/widower | 10,775 (20.3) |  |
| Charlson (%) |  |  |  |
|  | 0 | 39,256 (73.9) | 80,236 (64.0) |
|  | 1 | 8,117 (15.3) | 33,447 (26.7) |
|  | 2 | 3,762 ( 7.1) | 9,049 (7.2) |
|  | 3 | 1,176 ( 2.2) | 2,116 (1.7) |
|  | 4+ | 788 ( 1.5) | 580 (0.5) |
| Elixhauser (%) |  |  |  |
|  | 0 | 27,773 (52.3) | 33,546 (26.8) |
|  | 1 | 13,766 (25.9) | 42,467 (33.9) |
|  | 2 | 7,239 (13.6) | 28,188 (22.5) |
|  | 3+ | 4,321 ( 8.1) | 21,227 (16.9) |
| AIDS/HIV (%) |  | 5 ( 0.0) | 0 |
| Anaemia (%) |  | 416 ( 0.8) | 2,306 (1.8) |
| Arrhythmia (%) |  | 4,505 ( 8.5) | 13,040 (10.4) |
| Arterial hypertension (%) |  | 16,677 (31.4) | 66,837 (53.3) |
| Cancer (%) |  | 2,715 ( 5.1) | 3,055 (2.4) |
| CNS disease (%) |  | 1,682 ( 3.2) | 7,075 (5.6) |
| Coagulopathy (%) |  | 192 ( 0.4) | 583 (0.5) |
| Diabetes (%) |  | 4,077 ( 7.7) | 13,874 (11.1) |
| Drug alcohol abuse (%) |  | 223 ( 0.4) | 2,057 (1.6) |
| Fluid electrolyte disorders (%) |  | 304 ( 0.6) | 2,774 (2.2) |
| Heart condition (%) |  | 2,639 ( 5.0) | 5,746 (4.6) |
| Myocardial infarction (%) |  | 2,186 ( 4.1) | 3,774 (3.0) |
| Hypothyroidism (%) |  | 1,791 ( 3.4) | 10,453 (8.3) |
| Kidney disease (%) |  | 551 ( 1.0) | 7,370 (5.9) |
| Liver disease (%) |  | 207 ( 0.4) | 630 (0.5) |
| Lung airways disease (%) |  | 2,878 ( 5.4) | 18,337 (14.6) |
| Diagnosed obesity (%) |  | 1,000 ( 1.9) | 10,380 (8.3) |
| Peptic ulcer (%) |  | 341 ( 0.6) | 685 (0.6) |
| Rheumatic disease (%) |  | 1,922 ( 3.6) | 5,956 (4.8) |
| Vascular disease (%) |  | 1,686 ( 3.2) | 3,656 (2.9) |
| Weight loss (%) |  | 35 ( 0.1) | 509 (0.4) |

##### PAGE BREAK

Table 4: Variables selected by the bootstrap ranking procedure, and therefore kept in the main model. Only variables chosen at least 33 times out of 100 were kept in the reduced model. ASA classes II and III were, however, kept in both models. Variables chosen more frequently are likely to be more important predictors. (CNS = central nervous system.)

|  |  |
| --- | --- |
| variable | n |
| Cancer | 100 |
| CNS disease | 100 |
| Kidney disease | 100 |
| ASA class III | 100 |
| Age | 50 |
| Diagnosed obesity | 49 |
| Male sex | 38 |
| Heart condition | 20 |
| Anaemia | 9 |
| ASA class II | 8 |
| Myocardial infarction | 6 |

##### PAGE BREAK

Table 5: Estimated coefficients (beta) and odds ratios (OR) with 95 % confidence intervals for the reduced model.

|  |  |  |  |
| --- | --- | --- | --- |
| term | beta | OR 95 % CI | p |
| (intercept) | -13.28 |  |  |
| Cancer | 0.71 | 2.04 (1.30-3.08) | <0.01 |
| CNS disease | 0.90 | 2.45 (1.35-4.11) | <0.01 |
| Kidney disease | 1.33 | 3.79 (2.04-6.52) | <0.01 |
| ASA class II | 0.88 | 2.42 (1.20-5.78) | 0.03 |
| ASA class III | 1.66 | 5.27 (2.56-12.77) | <0.01 |
| Age | 0.08 | 1.08 (1.06-1.11) | <0.01 |
| Diagnosed obesity | 0.77 | 2.17 (0.91-4.36) | 0.05 |
| Male sex | 0.61 | 1.85 (1.36-2.52) | <0.01 |

##### PAGE BREAK

# Bibliography

1. **Garland A, Gordon M, Garellick G, Kärrholm J, Sköldenberg O, Hailer NP**. Risk of early mortality after cemented compared with cementless total hip arthroplasty. *The Bone & Joint Journal* 2017;99-B(1):37–43.

2. **Inacio MCS, Pratt NL, Roughead EE, Graves SE**. Using medications for prediction of revision after total joint arthroplasty. *The Journal of arthroplasty* 2015;30(12):2061–70.

3. **Gordon M, Stark A, Sköldenberg OG, Kärrholm J, Garellick G**. The influence of comorbidity scores on re-operations following primary total hip replacement: Comparison and validation of three comorbidity measures. *The Bone & Joint Journal* 2013;95-B(9):1184–91.

4. **Hofstede SN, Gademan MGJ, Vlieland TPMV, Nelissen RGHH, Mheen PJM-v de**. Preoperative predictors for outcomes after total hip replacement in patients with osteoarthritis: A systematic review. *BMC musculoskeletal disorders* 2016;17:212.

5. **Bülow E, Rolfson O, Cnudde P, Rogmark C, Garellick G, Nemes S**. Comorbidity does not predict long-term mortality after total hip arthroplasty. *Acta Orthopaedica* 2017;88(July):1–6.

6. **Bozic KJ, Ong K, Lau E, Berry DJ, Vail TP, Kurtz SM, et al.** Estimating risk in medicare patients with THA: An electronic risk calculator for periprosthetic joint infection and mortality. *Clinical Orthopaedics and Related Research®* 2013;471(2):574–583.

7. **Bozic KJ**. Orthopaedic healthcare worldwide: Shared medical decision making in orthopaedics. *Clinical Orthopaedics and Related Research* 2013;471(5):1412–1414.

8. **Harris AHS, Kuo AC, Bozic KJ, Lau E, Bowe T, Gupta S, et al.** American joint replacement registry risk calculator does not predict 90-day mortality in veterans undergoing total joint replacement. *Clinical Orthopaedics and Related Research* 2018;476(9):1869–1875.

9. **Harris AH, Kuo AC, Bowe T, Gupta S, Nordin D, Giori NJ**. Prediction models for 30-Day mortality and complications after total knee and hip arthroplasties for veteran health administration patients with osteoarthritis. *The Journal of Arthroplasty* 2018;33(5):1539–1545.

10. **Inacio M, Pratt N, Roughead E, Graves S**. Evaluation of three co-morbidity measures to predict mortality in patients undergoing total joint arthroplasty. *Osteoarthritis and cartilage* 2016;24(10):1718–1726.

11. **Price A, Smith J, Dakin H, Kang S, Eibich P, Cook J, et al.** The Arthroplasty Candidacy Help Engine tool to select candidates for hip and knee replacement surgery: Development and economic modelling. *Health Technology Assessment* 2019;23(32):1–216.

12. **Harris AHS, Kuo AC, Weng Y, Trickey AW, Bowe T, Giori NJ**. Can machine learning methods produce accurate and easy-to-use prediction models of 30-day complications and mortality after knee or hip arthroplasty? *Clinical Orthopaedics and Related Research* 2019;477(2):452–460.

13. **Manning DW, Edelstein AI, Alvi HM**. Risk prediction tools for hip and knee arthroplasty. *The Journal of the American Academy of Orthopaedic Surgeons* 2016;24(1):19–27.

14. **Bülow E, Nemes S, Rolfson O**. Are the first or the second hips of staged bilateral THAs more similar to unilateral procedures? A study from the swedish hip arthroplasty register. *Clinical Orthopaedics and Related Research* 2020;2020(478):11262–1270.

15. **Kärrholm J, Mohaddes M, Odin D, Vinblad J, Rogmark C, Rolfson O**. Svenska höftprotesregistret årsrapport 2017. 2018.

16. **Cnudde P, Rolfson O, Nemes S, Kärrholm J, Rehnberg C, Rogmark C, et al.** Linking Swedish health data registers to establish a research database and a shared decision-making tool in hip replacement. *BMC Musculoskeletal Disorders* 2016;17(1):414.

17. **Ludvigsson JF, Svedberg P, Olén O, Bruze G, Neovius M**. The longitudinal integrated database for health insurance and labour market studies (LISA) and its use in medical research. *European Journal of Epidemiology* 2019;34(4):423–437.

18. **Ludvigsson JF, Andersson E, Ekbom A, Feychting M, Kim J-L, Reuterwall C, et al.** External review and validation of the Swedish national inpatient register. *BMC Public Health* 2011;11(1):450.

19. **The NJR Editorial Board**. NJR 16th Annual Report 2019.Pdf.

20. **National Joint Registry (NJR)**. Data-Completeness-and-quality. (date last accessed 2 August 2020).

21. **Thorn JC, Turner E, Hounsome L, Walsh E, Donovan JL, Verne J, et al.** Validation of the hospital episode statistics outpatient dataset in england. *PharmacoEconomics* 2016;34(2):161–168.

22. **Sayers A, Ben-Shlomo Y, Blom AW, Steele F**. Probabilistic record linkage. *International Journal of Epidemiology* 2016;45(3):954–964.

23. **Quan H, Sundararajan V, Halfon P, Fong A, Burnand B, Luthi J-C, et al.** Coding algorithms for defining comorbidities in ICD-9-CM and ICD-10 administrative data. *Medical care* 2005;43(11):1130–1139.

24. **Guo P, Zeng F, Hu X, Zhang D, Zhu S, Deng Y, et al.** Improved variable selection algorithm using a LASSO-Type penalty, with an application to assessing hepatitis B infection relevant factors in community residents. Emmert-Streib F, ed. *PLOS ONE* 2015;10(7):e0134151.

25. **Miller ME, Hui SL, Tierney WM**. Validation techniques for logistic regression models. *Statistics in Medicine* 1991;10(8):1213–1226.

26. **Nattino G, Finazzi S, Bertolini G**. A new test and graphical tool to assess the goodness of fit of logistic regression models. *Statistics in Medicine* 2016;35(5):709–720.

27. **Steyerberg EW, Borsboom GJJM, Houwelingen HC van, Eijkemans MJC, Habbema JDF**. Validation and updating of predictive logistic regression models: A study on sample size and shrinkage. *Statistics in Medicine* 2004;23(16):2567–2586.

28. **Gibbs VN, McCulloch RA, Dhiman P, McGill A, Taylor AH, Palmer AJR, et al.** Modifiable risk factors for mortality in revision total hip arthroplasty for periprosthetic fracture. *The Bone & Joint Journal* 2020 [cited 24 Aug 2020];102-B(5):580–585.

29. **Burns EM, Rigby E, Mamidanna R, Bottle A, Aylin P, Ziprin P, et al.** Systematic review of discharge coding accuracy. *Journal of Public Health* 2012;34(1):138–148.

30. **Smith AJ, Dieppe P, Porter M, Blom AW**. Risk of cancer in first seven years after metal-on-metal hip replacement compared with other bearings and general population: Linkage study between the National Joint Registry of England and Wales and hospital episode statistics. *BMJ (Online)* 2012;344.

31. **Mouchti S, Whitehouse MR, Sayers A, Hunt LP, MacGregor A, Blom AW**. The association of body mass index with risk of long-term revision and 90-day mortality following primary total hip replacement. *The Journal of Bone and Joint Surgery* 2018;100(24):2140–2152.

32. **Cnudde P**. Longitudinal outcomes following total hip replacement: Time trends, sequence of events and study of factors influencing implant survival and mortality [PhD thesis]. 2018. [cited 27 Aug 2020].

33. **Bülow E**. Predicting mortality by comorbidity for patients with hip arthroplasty: Prospective observational register studies of a nationwide Swedish cohort [PhD thesis]. 2020. [cited 27 Aug 2020].